

SOCY7708: Hierarchical Linear Modeling
Instructor: Natasha Sarkisian
Class notes: Interpreting Interactions in HLM

So far, we looked at a cross-level interaction of a continuous level-1 variable, SES, and a dichotomous level 2 variable, SECTOR, but there can be various other combinations; we will discuss four:

- (1) Two dichotomous variables
- (2) A dichotomy and a continuous variable
- (3) A multicategory variable and a continuous variable
- (4) Two continuous variables

What is crucial for interpretation in all of these cases is to consider the size and significance of what's called simple slopes – that is, the slope of X when moderator Z is at a specified value. That means, we should always consider one variable as the focal one (X) and one as the moderator (Z). It's arbitrary for a given interaction which one is which – but it usually makes substantive sense to designate one variable as the focal one and the other as the moderator. In HLM, we typically discuss explaining variance in level 1 slopes using level 2 variables, so that would suggest that level 1 variable would serve as your focal variable, and level 2 variable serve as a moderator, but we can always invert that interpretation – for example, rather than discussing how the effect of SES is different in Catholic vs public schools (ses=focal, sector=moderator), we could discuss the size of the school type effect for students with different levels of SES – for instance, high, medium, low (sector=focal, ses=moderator). In what follows, I will discuss interpretation for each of the four possible variable combinations and show some tools for easier interpretation.

Example 1: Two dichotomous variables

Here, we will look at a cross-level interaction of sector on level 2 and minority on level 1:

```
. mixed mathach i.minority##i.sector ses size || id:minority, cov(unstr)
```

Mixed-effects ML regression
Group variable: id

Number of obs = 7,185
Number of groups = 160
Obs per group:
min = 14
avg = 44.9
max = 67

Wald chi2(5) = 879.58
Prob > chi2 = 0.0000

Log likelihood = -23175.932

mathach	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
1.minority	-4.226579	.3040739	-13.90	0.000	-4.822553	-3.630605
1.sector	2.266543	.3311193	6.85	0.000	1.617561	2.915525
minority#sector						
1 1	2.187448	.4225161	5.18	0.000	1.359331	3.015564
ses	2.04651	.1056623	19.37	0.000	1.839416	2.253605
size	.0009792	.0002577	3.80	0.000	.000474	.0014844
_cons	11.44099	.3914482	29.23	0.000	10.67376	12.20821

```
-----
Random-effects parameters | Estimate Std. err. [95% conf. interval]
-----+-----
id: Unstructured         |
    var(minority)        |   .5397437   .6465572   .0515865   5.647276
    var(_cons)           |   2.319074   .4180191   1.628858   3.301763
    cov(minority,_cons)  |  -.0819745   .4043974  -.8745789   .7106299
-----+-----
    var(Residual)       |   35.93477   .6106514   34.75763   37.15178
-----
LR test vs. linear model: chi2(3) = 205.32          Prob > chi2 = 0.0000
```

Note: LR test is conservative and provided only for reference.

Minority as the focal variable, sector as the moderator:

For public schools, minority students have math achievement scores that are 4.2 units lower than non-minority students. For Catholic schools, minority students have math achievement scores that are 2 units lower ($-4.2+2.2=-2$) than non-minority students. So the effect of being a minority is more pronounced in public than in Catholic schools, and difference in effects is statistically significant (as indicated by the significant interaction term). We know that the effect of being a minority is significant for public schools, but what about in Catholic schools? We can find out if we reverse the dummy variable (by running the same mixed command with `ib1.minority`) or using margins command:

```
. margins, dydx(minority) at(sector=(0 1)) atmeans

Conditional marginal effects          Number of obs = 7,185

Expression: Linear prediction, fixed portion, predict()
dy/dx wrt:  1.minority
1._at: 0.minority = .725261 (mean)
      1.minority = .274739 (mean)
      sector    = 0
      ses       = .0001434 (mean)
      size      = 1056.862 (mean)
2._at: 0.minority = .725261 (mean)
      1.minority = .274739 (mean)
      sector    = 1
      ses       = .0001434 (mean)
      size      = 1056.862 (mean)
```

```
-----
|              Delta-method
|              dy/dx  std. err.      z    P>|z|    [95% conf. interval]
-----+-----
0.minority | (base outcome)
-----+-----
1.minority |
  _at |
  1 | -4.226579   .3040739  -13.90   0.000   -4.822553   -3.630605
  2 | -2.039132   .296945   -6.87    0.000   -2.621133   -1.45713
-----
```

Note: dy/dx for factor levels is the discrete change from the base level.

Here, we can see that even though the effect of being a minority is smaller in Catholic schools than in public schools, it is nevertheless significant in both types of schools.

Sector as the focal variable, minority as the moderator:

Among non-minority students, those in Catholic schools have math achievement scores that are 2.3 units higher than those in public schools. And among minority students, that gap is even higher – minority students in Catholic schools have scores that are 4.5 units higher than minority students in public schools. Again, this difference in effects of school type is significant because the interaction term is significant. And margins command will help us see if simple slopes of sector are significant at both values of minority variable:

```
. margins, dydx(sector) at(minority=(0 1)) atmeans
```

Conditional marginal effects Number of obs = 7,185

Expression: Linear prediction, fixed portion, predict()
dy/dx wrt: 1.sector

```
1._at: minority = 0
      0.sector = .5068894 (mean)
      1.sector = .4931106 (mean)
      ses      = .0001434 (mean)
      size     = 1056.862 (mean)
2._at: minority = 1
      0.sector = .5068894 (mean)
      1.sector = .4931106 (mean)
      ses      = .0001434 (mean)
      size     = 1056.862 (mean)
```

		Delta-method				[95% conf. interval]	
		dy/dx	std. err.	z	P> z		
0.sector		(base outcome)					
1.sector							
	_at						
	1	2.266543	.3311193	6.85	0.000	1.617561	2.915525
	2	4.45399	.4630216	9.62	0.000	3.546485	5.361496

Note: dy/dx for factor levels is the discrete change from the base level.

Here, we can see that indeed, the effect of school type is significant for both minority and non-minority students. Finally, to further assist our interpretation, we could calculate predicted values of math achievement for the 4 groups based on minority and sector variables:

```
. margins, at(minority=(0 1) sector=(0 1)) atmeans
```

Adjusted predictions Number of obs = 7,185

Expression: Linear prediction, fixed portion, predict()

```
1._at: minority = 0
      sector    = 0
      ses       = .0001434 (mean)
      size      = 1056.862 (mean)
2._at: minority = 0
      sector    = 1
      ses       = .0001434 (mean)
      size      = 1056.862 (mean)
3._at: minority = 1
      sector    = 0
      ses       = .0001434 (mean)
```

```

size      = 1056.862 (mean)
4._at: minority = 1
sector    = 1
ses       = .0001434 (mean)
size      = 1056.862 (mean)

```

	Margin	Delta-method std. err.	z	P> z	[95% conf. interval]	
_at						
1	12.47616	.213089	58.55	0.000	12.05851	12.8938
2	14.7427	.2364074	62.36	0.000	14.27935	15.20605
3	8.249577	.3181134	25.93	0.000	7.626086	8.873068
4	12.70357	.3115603	40.77	0.000	12.09292	13.31421

The scores are highest for non-minority students in Catholic schools and lowest for minority students in public schools.

Example 2: A dichotomy and a continuous variable

Here, we will look at a dichotomous variable minority on level 1 and a continuous variable size on level 2, but similar interpretation approaches can be used if your continuous variable is on level 1 and your dichotomy is on level 2.

```
. sum size
```

Variable	Obs	Mean	Std. dev.	Min	Max
size	7,185	1056.862	604.1725	100	2713

```
. egen tagged=tag(id)
```

```
. sum size if tagged==1
```

Variable	Obs	Mean	Std. dev.	Min	Max
size	160	1097.825	629.5064	100	2713

```
. gen sizem=size-r(mean)
```

```
. mixed mathach i.minority##c.sizem || id: minority, cov(unstr)
```

```

Mixed-effects ML regression      Number of obs   =      7,185
Group variable: id              Number of groups =      160
                                Obs per group:
                                min =      14
                                avg =     44.9
                                max =      67
                                Wald chi2(3)      =     224.93
                                Prob > chi2       =      0.0000
Log likelihood = -23393.211

```

	mathach	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
1.minority		-3.723111	.2577542	-14.44	0.000	-4.2283	-3.217922
size		.0002325	.0003452	0.67	0.501	-.0004442	.0009092
minority#c.sizem							
1		-.0015012	.0004082	-3.68	0.000	-.0023012	-.0007012
_cons		13.69733	.209928	65.25	0.000	13.28588	14.10878

```
-----
```

Random-effects parameters	Estimate	Std. err.	[95% conf. interval]	

id: Unstructured				
var(minority)	2.262594	.9714639	.9752954	5.249008
var(_cons)	5.561104	.78949	4.210357	7.34519
cov(minority,_cons)	.9227327	.6520362	-.3552349	2.2007

var(Residual)	37.41324	.6361601	36.18693	38.6811

LR test vs. linear model: chi2(3) = 741.52 Prob > chi2 = 0.0000

Note: LR test is conservative and provided only for reference.

Minority as the focal variable, size as the moderator:

In schools of average (mean) size, minority students have math achievement scores that are 3.7 units lower than non-minority students. If school size is one SD (629.5) higher than average, minority students have math achievement scores that are 4.6 units lower than non-minority students (calculated as $-3.7 - .0015012 * 629.5$).

```
. di -3.7 - .0015012*629.5
-4.6450054
```

If school size is one SD lower than average, minority students have math achievement scores that are 2.8 units lower than non-minority students.

```
. di -3.7 + .0015012*629.5
-2.7549946
```

So we conclude that the race/ethnicity gap is more pronounced in larger schools than in smaller schools, and that difference is statistically significant (as indicated by the significant interaction term). To see if simple slopes of minority are significant at various levels of size variable, we employ margins command:

```
. sum sizem if tagged==1
```

Variable	Obs	Mean	Std. dev.	Min	Max
sizem	160	-3.86e-06	629.5064	-997.825	1615.175

```
. return list
```

scalars:

```

r(N) = 160
r(sum_w) = 160
r(mean) = -3.86238098145e-06
r(Var) = 396278.3592549189
r(sd) = 629.5064409955778
r(min) = -997.8250122070313
r(max) = 1615.175048828125
r(sum) = -.0006179809570313
```

```
. global min=r(min)
. global max=r(max)
```

```
. global mean=r(mean)
. global plusd=r(mean)+r(sd)
. global minusd=r(mean)-r(sd)
. margins, dydx(minority) at(size=($min $minusd $mean $plusd $max)) atmeans
```

Conditional marginal effects Number of obs = 7,185

Expression: Linear prediction, fixed portion, predict()

```
dy/dx wrt: 1.minority
1._at: 0.minority = .725261 (mean)
       1.minority = .274739 (mean)
       sizem      = -997.825
2._at: 0.minority = .725261 (mean)
       1.minority = .274739 (mean)
       sizem      = -629.5064
3._at: 0.minority = .725261 (mean)
       1.minority = .274739 (mean)
       sizem      = -3.86e-06
4._at: 0.minority = .725261 (mean)
       1.minority = .274739 (mean)
       sizem      = 629.5064
5._at: 0.minority = .725261 (mean)
       1.minority = .274739 (mean)
       sizem      = 1615.175
```

		Delta-method				[95% conf. interval]	
		dy/dx	std. err.	z	P> z		

0.minority		(base outcome)					

1.minority							
	_at						
	1	-2.225132	.4869101	-4.57	0.000	-3.179458	-1.270805
	2	-2.778068	.3680558	-7.55	0.000	-3.499444	-2.056692
	3	-3.723111	.2577542	-14.44	0.000	-4.2283	-3.217922
	4	-4.668154	.359792	-12.97	0.000	-5.373333	-3.962974
	5	-6.147883	.7023876	-8.75	0.000	-7.524537	-4.771229

Note: dy/dx for factor levels is the discrete change from the base level.

Size as the focal variable, minority as the moderator:

For non-minority students, if a school has one extra student (school size increases by 1), then their math achievement increases by .0002 of a unit, but that increase is not statistically significant. For minority students, if the school size increases by 1, their math achievement decreases by .001 of a unit:

```
. di .0002325-.0015012
-.0012687
```

But we don't know yet if that's a significant decrease – we need to use margins to look at that simple slope:

```
. margins, dydx(sizem) at(minority=(0 1)) atmeans
```

Conditional marginal effects Number of obs = 7,185

```

Expression: Linear prediction, fixed portion, predict()
dy/dx wrt:  size
1._at: minority = 0
      size      = -40.96321 (mean)
2._at: minority = 1
      size      = -40.96321 (mean)

```

		Delta-method				
		dy/dx	std. err.	z	P> z	[95% conf. interval]
size	_at					
	1	.0002325	.0003452	0.67	0.501	-.0004442 .0009092
	2	-.0012687	.0004836	-2.62	0.009	-.0022166 -.0003209

Here, we can see that the effect of school size is statistically significant for minority students but not for the non-minority students. The size of coefficient for minority students still looks small, but that's because one unit for school size is small – it's one student. Let's look how much change that would be for one SD increase in school size:

```

. di -.0012687*629.5
-.79864665

```

We can also show these slopes on a graph, but to have a better scale, we will reestimate the model with size being uncentered (the results don't substantively change from that although the constant and main effects coefficients are less interpretable, but the graph is exactly the same, just labeled better):

```

. mixed mathach i.minority#c.size || id: minority, cov(unstr)

```

```

Mixed-effects ML regression      Number of obs   =      7,185
Group variable: id              Number of groups =      160
                                Obs per group:
                                min =      14
                                avg =     44.9
                                max =      67
                                Wald chi2(3)      =     224.93
                                Prob > chi2       =     0.0000
Log likelihood = -23393.211

```

		Coefficient	Std. err.	z	P> z	[95% conf. interval]
1.minority	size	-2.075007	.5219894	-3.98	0.000	-3.098088 -1.051927
		.0002325	.0003452	0.67	0.501	-.0004442 .0009092
minority#c.size	1	-.0015012	.0004082	-3.68	0.000	-.0023012 -.0007012
	_cons	13.44209	.4274142	31.45	0.000	12.60437 14.2798

Random-effects parameters		Estimate	Std. err.	[95% conf. interval]	
id: Unstructured					
	var(minority)	2.262594	.9714639	.9752954	5.249008
	var(_cons)	5.561104	.78949	4.210357	7.34519
	cov(minority,_cons)	.9227327	.6520362	-.3552349	2.2007
	var(Residual)	37.41324	.6361601	36.18693	38.6811

 LR test vs. linear model: chi2(3) = 741.52 Prob > chi2 = 0.0000

Note: LR test is conservative and provided only for reference.

. sum size if tagged==1

Variable	Obs	Mean	Std. dev.	Min	Max
size	160	1097.825	629.5064	100	2713

. global min=r(min)

. global max=r(max)

. global mean=r(mean)

. global plusd=r(mean)+r(sd)

. global minusd=r(mean)-r(sd)

. margins, at(minority=(0 1) size=(\$min \$minusd \$mean \$plusd \$max))

Adjusted predictions

Number of obs = 7,185

Expression: Linear prediction, fixed portion, predict()

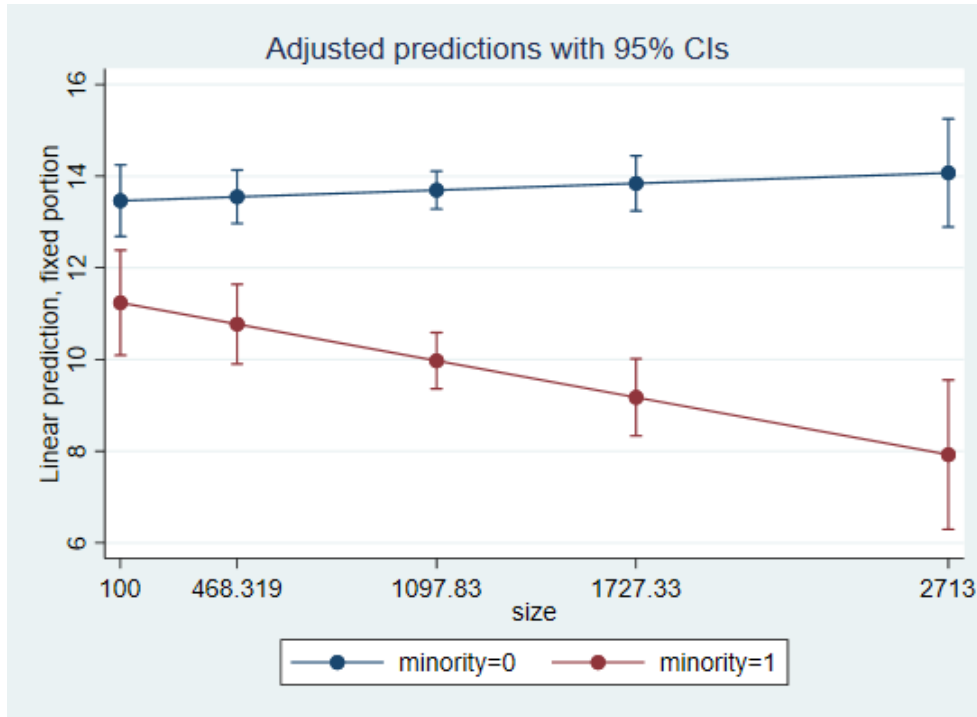
```

1._at: minority =      0
      size      =     100
2._at: minority =      0
      size      = 468.3186
3._at: minority =      0
      size      = 1097.825
4._at: minority =      0
      size      = 1727.331
5._at: minority =      0
      size      =     2713
6._at: minority =      1
      size      =     100
7._at: minority =      1
      size      = 468.3186
8._at: minority =      1
      size      = 1097.825
9._at: minority =      1
      size      = 1727.331
10._at: minority =      1
      size      =     2713
  
```

	Margin	Delta-method std. err.	z	P> z	[95% conf. interval]	
_at						
1	13.46534	.3976975	33.86	0.000	12.68586	14.24481
2	13.55097	.297343	45.57	0.000	12.96819	14.13375
3	13.69733	.209928	65.25	0.000	13.28588	14.10878
4	13.84368	.3069108	45.11	0.000	13.24215	14.44522
5	14.07285	.6020278	23.38	0.000	12.89289	15.2528
6	11.2402	.5838319	19.25	0.000	10.09592	12.38449
7	10.7729	.4436979	24.28	0.000	9.903269	11.64253
8	9.974216	.3125363	31.91	0.000	9.361656	10.58678
9	9.17553	.4287868	21.40	0.000	8.335124	10.01594
10	7.924965	.8313537	9.53	0.000	6.295541	9.554388

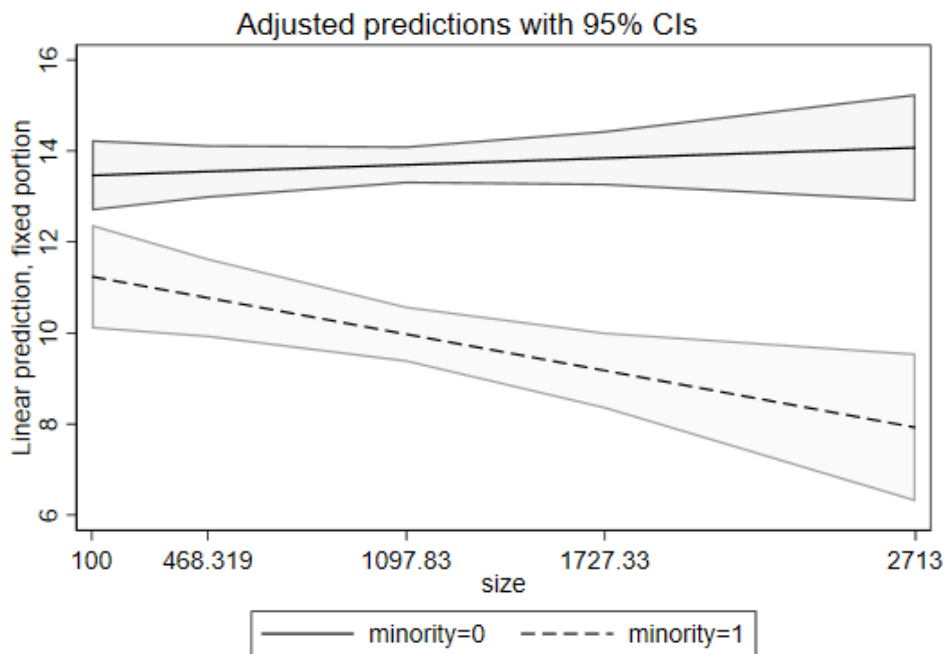

```
. marginsplot, x(size)
```

Variables that uniquely identify margins: minority size



```
. marginsplot, x(size) plotop(msymbol(i)) recastci(rarea) ciopts(fintensity(5))  
scheme(slmono)
```

Variables that uniquely identify margins: minority size



We can also graph the size of minority gap at different levels of school size – this graph would be used to assist interpretation when minority is the focal variable and school size is the moderator:

```
. margins, dydx(minority) at(size=($min $minussd $mean $plussd $max)) atmeans
Conditional marginal effects                                Number of obs = 7,185
```

Expression: Linear prediction, fixed portion, predict()
dy/dx wrt: 1.minority

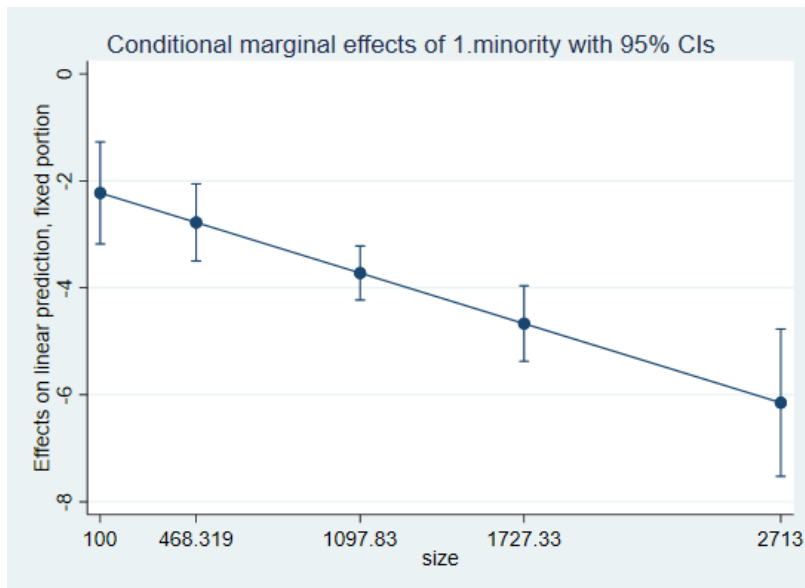
```
1._at: 0.minority = .725261 (mean)
        1.minority = .274739 (mean)
        size      = 100
2._at: 0.minority = .725261 (mean)
        1.minority = .274739 (mean)
        size      = 468.3186
3._at: 0.minority = .725261 (mean)
        1.minority = .274739 (mean)
        size      = 1097.825
4._at: 0.minority = .725261 (mean)
        1.minority = .274739 (mean)
        size      = 1727.331
5._at: 0.minority = .725261 (mean)
        1.minority = .274739 (mean)
        size      = 2713
```

		Delta-method				
		dy/dx	std. err.	z	P> z	[95% conf. interval]
0.minority		(base outcome)				
1.minority						
	_at					
	1	-2.225132	.4869101	-4.57	0.000	-3.179458 -1.270805
	2	-2.778068	.3680557	-7.55	0.000	-3.499444 -2.056692
	3	-3.723111	.2577542	-14.44	0.000	-4.2283 -3.217922
	4	-4.668154	.359792	-12.97	0.000	-5.373333 -3.962974
	5	-6.147883	.7023876	-8.75	0.000	-7.524537 -4.771229

Note: dy/dx for factor levels is the discrete change from the base level.

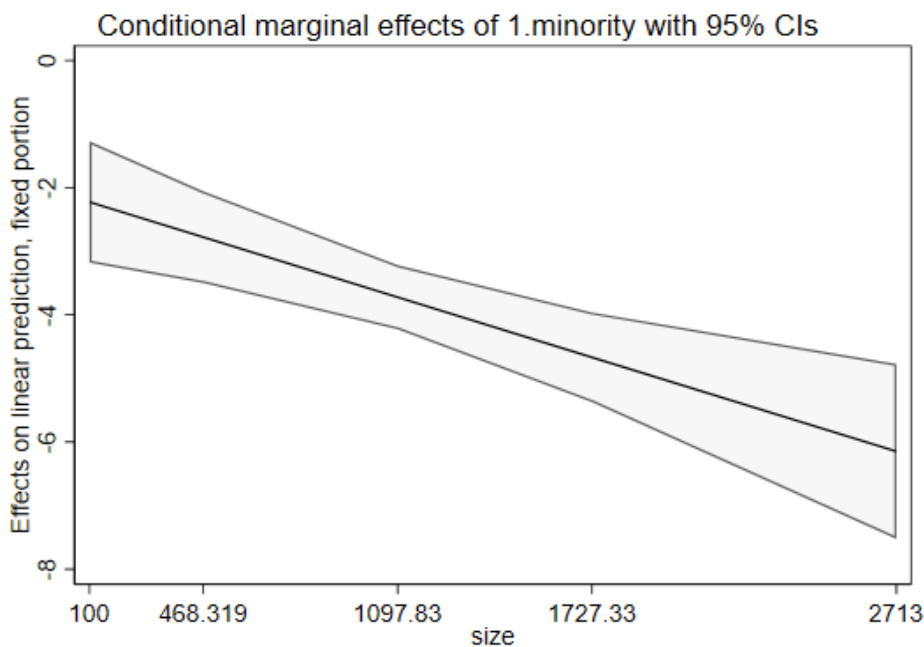
```
. marginsplot
```

Variables that uniquely identify margins: size



```
. marginsplot, plotop(msymbol(i)) recastci(rarea) ciopts(fintensity(5)) scheme(slmono)
```

Variables that uniquely identify margins: size



Example 3: A multcategory variable and a continuous variable

For this example, I will again create dummy variables for school size but this time, I'll create four of them, approximately based on quartiles. I will use a function of egen variable to create that. My level 1 continuous variable will be SES.

```
. egen sized=cut(size), group(4)
```

```
. tab sized
```

sized	Freq.	Percent	Cum.
-------	-------	---------	------

0	1,796	25.00	25.00
1	1,764	24.55	49.55
2	1,803	25.09	74.64
3	1,822	25.36	100.00
Total	7,185	100.00	

```
. mixed mathach c.ses##i.sized || id:ses, cov(unstr)
```

```
Mixed-effects ML regression      Number of obs   =      7,185
Group variable: id              Number of groups =      160
                                Obs per group:
                                min =      14
                                avg =     44.9
                                max =      67
                                Wald chi2(7)      =     457.54
                                Prob > chi2       =      0.0000
Log likelihood = -23308.899
```

mathach	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
ses	2.100147	.2363191	8.89	0.000	1.63697	2.563324
sized						
1	1.065607	.5351867	1.99	0.046	.01666	2.114553
2	.9430464	.5394171	1.75	0.080	-.1141917	2.000284
3	-.2007692	.5149239	-0.39	0.697	-1.210001	.8084631
sized#c.ses						
1	-.0825774	.3329181	-0.25	0.804	-.735085	.5699301
2	.3169989	.3335864	0.95	0.342	-.3368185	.9708163
3	.8111932	.3194897	2.54	0.011	.185005	1.437381
_cons	12.25458	.3819801	32.08	0.000	11.50592	13.00325

Random-effects parameters	Estimate	Std. err.	[95% conf. interval]	
id: Unstructured				
var(ses)	.3036592	.218077	.0743163	1.240762
var(_cons)	4.534295	.635098	3.445767	5.966693
cov(ses,_cons)	-.0604166	.2803227	-.6098389	.4890058
var(Residual)	36.80666	.6285147	35.59518	38.05937

```
LR test vs. linear model: chi2(3) = 428.23      Prob > chi2 = 0.0000
```

Note: LR test is conservative and provided only for reference.

First, I will test if the interaction terms between SES and SIZE dummies are jointly significant – you cannot judge their significance by individual coefficient significance tests because those will change depending on the omitted category.

```
. mat list e(b)
```

```
e(b) [1,14]
mathach:      mathach:      mathach:      mathach:      mathach:      mathach:
mathach:
              0b.              1.              2.              3.              0b.sized#
1.sized#
```

```

      ses      sized      sized      sized      sized      co.ses
c.ses
y1  2.1001473      0  1.0656066  .94304639  -.20076917      0  -
.08257741

      mathach:      mathach:      mathach:      lns1_1_1:      lns1_1_2:      atr1_1_1_2:
lnsig_e:
      2.sized#      3.sized#
      c.ses      c.ses      _cons      _cons      _cons      _cons
_cons
y1  .31699893  .81119323  12.254582  -.59592467  .75583481  -.05153382
1.8028394

. test 1.sized#c.ses=0

( 1) [mathach]1.sized#c.ses = 0

      chi2( 1) =      0.06
      Prob > chi2 =      0.8041

. test 2.sized#c.ses=0, acc

( 1) [mathach]1.sized#c.ses = 0
( 2) [mathach]2.sized#c.ses = 0

      chi2( 2) =      1.61
      Prob > chi2 =      0.4475

. test 3.sized#c.ses=0, acc

( 1) [mathach]1.sized#c.ses = 0
( 2) [mathach]2.sized#c.ses = 0
( 3) [mathach]3.sized#c.ses = 0

      chi2( 3) =      9.92
      Prob > chi2 =      0.0193

```

They are jointly significant. And in fact, if we wanted to better highlight that with significant coefficients when presenting the results, we may want to omit the last rather than the first SIZE category (so compare everything to the largest schools):

```

. mixed mathach c.ses##ib3.sized || id:ses, cov(unstr)

Computing standard errors ...

Mixed-effects ML regression      Number of obs      =      7,185
Group variable: id              Number of groups    =      160
                                Obs per group:
                                min =      14
                                avg =      44.9
                                max =      67
                                Wald chi2(7)      =      457.54
                                Prob > chi2        =      0.0000

Log likelihood = -23308.899

```

mathach	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
ses	2.911341	.2150045	13.54	0.000	2.489939	3.332742
sized						
0	.2007692	.5149239	0.39	0.697	-.8084631	1.210001
1	1.266376	.5096605	2.48	0.013	.2674596	2.265292

I may also want to estimate and present this model with an alternative parametrization that allows me to right away see SES slopes separately for each size category; this model doesn't include the interaction terms so we couldn't see if SES slopes are different depending on SIZE, so it should only be used as a follow-up to a standard model with interactions if we find that they are significant. This parametrization is called the separate slope parameterization, it includes four separate SES coefficients for the four SIZE groups rather than the more conventional interaction term parameterization (main effect of SES and three interaction terms for SIZE). For that, we create four variables in which one size-based group's SES values are included along with zeroes for the other three groups. These variables allow us to obtain separate simple slopes for SES for each size group. For more details on that approach, or to justify using it, see: Cohen, Jacob, Patricia Cohen, Stephen G. West, and Leona Aiken. 2003. *Applied Multiple Regression/Correlation Analysis for the Behavioral Sciences*, 3rd ed. Mahwah, NJ: Lawrence Erlbaum.

```
. tab sized, gen(sized_)

      sized |      Freq.   Percent   Cum.
-----+-----
      0 |      1,796    25.00    25.00
      1 |      1,764    24.55    49.55
      2 |      1,803    25.09    74.64
      3 |      1,822    25.36   100.00
-----+-----
    Total |      7,185   100.00

. for num 1/4: gen ses_sized_X=ses*sized_X
-> gen ses_sized_1=ses*sized_1
-> gen ses_sized_2=ses*sized_2
-> gen ses_sized_3=ses*sized_3
-> gen ses_sized_4=ses*sized_4

. mixed mathach i.sized ses_sized* || id:ses, cov(unstr)

Mixed-effects ML regression              Number of obs   =      7,185
Group variable: id                       Number of groups =      160
                                           Obs per group:
                                           min =          14
                                           avg =         44.9
                                           max =           67
                                           Wald chi2(7)    =      457.54
                                           Prob > chi2     =      0.0000

Log likelihood = -23308.899
-----+-----
      mathach | Coefficient  Std. err.      z    P>|z|    [95% conf. interval]
-----+-----
      sized |
      1 |      1.065607  .5351867     1.99  0.046     .01666     2.114553
      2 |      .9430464  .5394171     1.75  0.080    -.1141917     2.000284
      3 |     -.2007692  .5149239    -0.39  0.697    -1.210001     .8084631
      ses_sized_1 |      2.100147  .2363191     8.89  0.000     1.63697     2.563324
      ses_sized_2 |      2.01757   .2344947     8.60  0.000     1.557969     2.477171
      ses_sized_3 |      2.417146  .2354425    10.27  0.000     1.955687     2.878605
      ses_sized_4 |      2.911341  .2150045    13.54  0.000     2.489939     3.332742
```

```

      _cons | 12.25458 .3819801 32.08 0.000 11.50592 13.00325
-----+-----
Random-effects parameters | Estimate Std. err. [95% conf. interval]
-----+-----
id: Unstructured
      var(ses) | .3036591 .2180771 .0743162 1.240763
      var(_cons) | 4.534295 .635098 3.445767 5.966693
      cov(ses,_cons) | -.0604166 .2803227 -.6098389 .4890057
-----+-----
      var(Residual) | 36.80666 .6285147 35.59518 38.05937
-----+-----
LR test vs. linear model: chi2(3) = 428.23 Prob > chi2 = 0.0000

```

Note: LR test is conservative and provided only for reference.

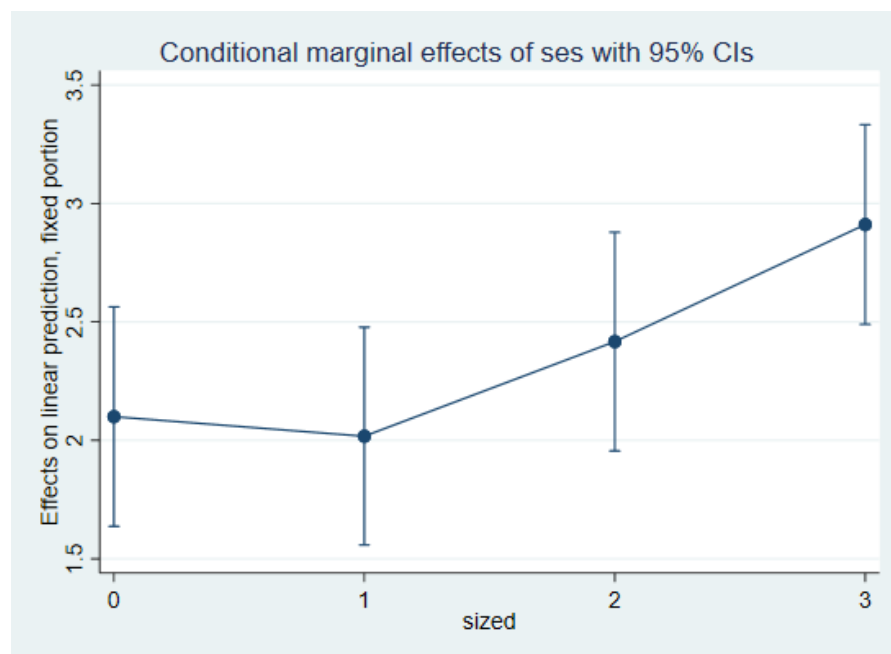
I omitted the smallest school category, but perhaps omitting the largest could be better; that doesn't affect the simple slopes for SES that we can see in this model, however.

Next, we turn to graphic representation of our interactions results. Since my SIZED variable is ordinal, I could consider a graph illustrating the size of SES effect for each SIZE category:

```

. marginsplot
Variables that uniquely identify margins: sized

```



We could also illustrate this with a graph showing actual slopes of SES at each level of SIZE – for that, we do margins without dydx to generate predicted values:

```

. sum ses

```

Variable	Obs	Mean	Std. dev.	Min	Max
ses	7,185	.0001434	.7793552	-3.758	2.692


```

. global min=r(min)
. global max=r(max)
. global mean=r(mean)
. global plusd=r(mean)+r(sd)
. global minusd=r(mean)-r(sd)
. margins, at(sized=(0 1 2 3) ses=($min $minusd $mean $plusd $max)) atmeans

```

Adjusted predictions

Number of obs = 7,185

Expression: Linear prediction, fixed portion, predict()

```

1._at: ses = -3.758
      sized = 0
2._at: ses = -3.758
      sized = 1
3._at: ses = -3.758
      sized = 2
4._at: ses = -3.758
      sized = 3
5._at: ses = -.7792118
      sized = 0
6._at: ses = -.7792118
      sized = 1
7._at: ses = -.7792118
      sized = 2
8._at: ses = -.7792118
      sized = 3
9._at: ses = .0001434
      sized = 0
10._at: ses = .0001434
      sized = 1
11._at: ses = .0001434
      sized = 2
12._at: ses = .0001434
      sized = 3
13._at: ses = .7794985
      sized = 0
14._at: ses = .7794985
      sized = 1
15._at: ses = .7794985
      sized = 2
16._at: ses = .7794985
      sized = 3
17._at: ses = 2.692
      sized = 0
18._at: ses = 2.692
      sized = 1
19._at: ses = 2.692
      sized = 2
20._at: ses = 2.692
      sized = 3

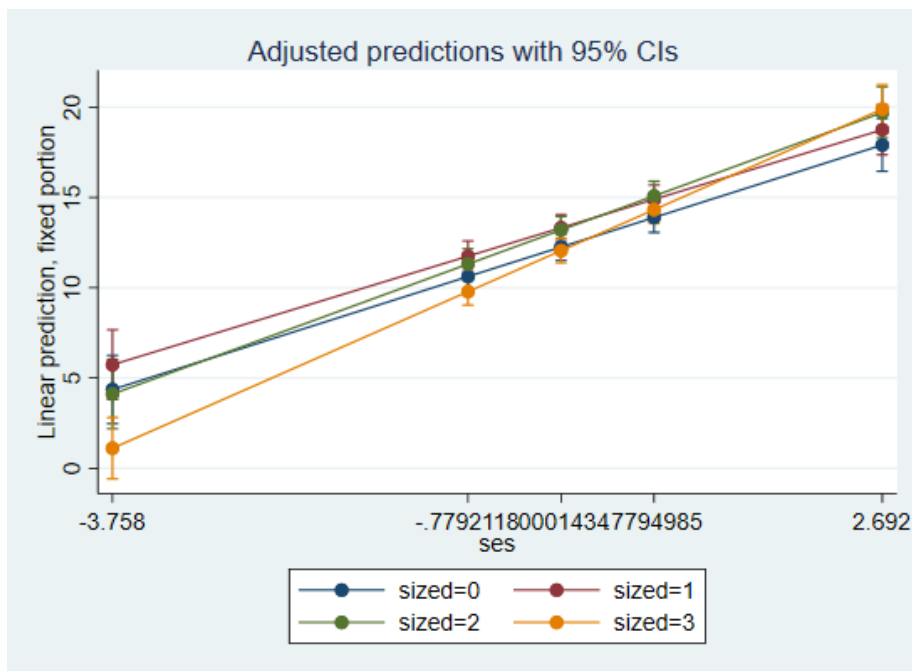
```

		Delta-method				
		Margin	std. err.	z	P> z	[95% conf. interval]
-----		-----				
_at						
1		4.362229	.9643754	4.52	0.000	2.472088 6.25237
2		5.738162	.984626	5.83	0.000	3.80833 7.667993

3		4.113994	.97967	4.20	0.000	2.193876	6.034112
4		1.112996	.8670579	1.28	0.199	-.5864064	2.812398
5		10.61812	.4229256	25.11	0.000	9.789204	11.44704
6		11.74807	.4298475	27.33	0.000	10.90559	12.59056
7		11.31416	.4304879	26.28	0.000	10.47042	12.1579
8		9.785262	.3782832	25.87	0.000	9.043841	10.52668
9		12.25488	.3819804	32.08	0.000	11.50622	13.00355
10		13.32048	.3748519	35.54	0.000	12.58578	14.05517
11		13.19798	.3808684	34.65	0.000	12.45149	13.94446
12		12.05423	.3453094	34.91	0.000	11.37744	12.73102
13		13.89164	.4251986	32.67	0.000	13.05827	14.72502
14		14.89288	.403804	36.88	0.000	14.10144	15.68432
15		15.08179	.4148992	36.35	0.000	14.2686	15.89498
16		14.3232	.389275	36.79	0.000	13.56023	15.08616
17		17.90818	.7442504	24.06	0.000	16.44947	19.36688
18		18.75149	.7081423	26.48	0.000	17.36355	20.13942
19		19.70459	.7238617	27.22	0.000	18.28584	21.12333
20		19.89114	.6846703	29.05	0.000	18.54921	21.23307

```
. marginsplot, x(ses)
```

Variables that uniquely identify margins: sized ses

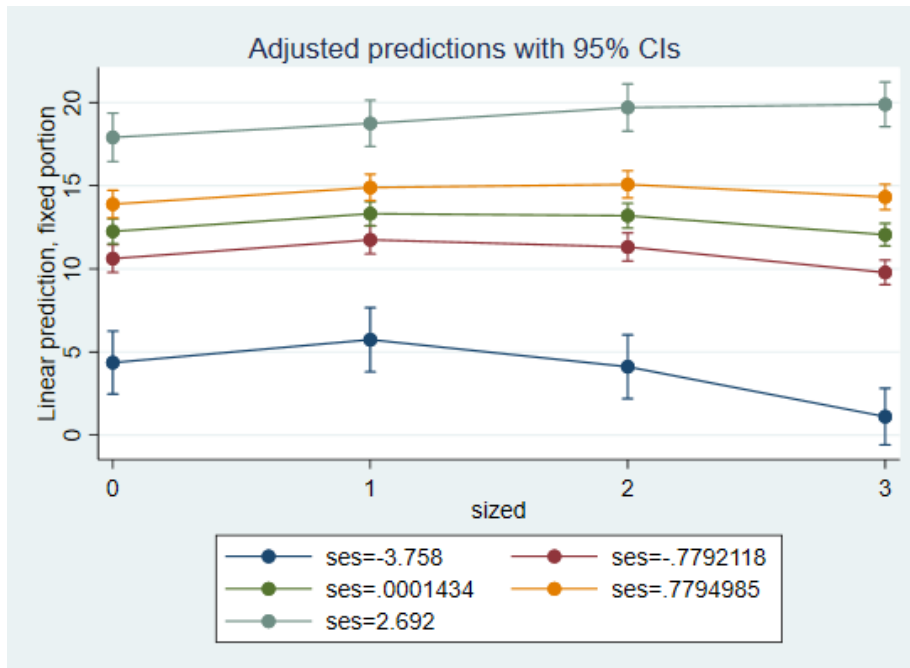


SIZE as focal, SES as a moderator:

We can use the same predicted values to create a graph focusing on SIZE as focal:

```
. marginsplot, x(sized)
```

Variables that uniquely identify margins: sized ses



To examine which SIZE effects are statistically significant, we do margins with dydx focusing on sized:

```
. margins, dydx(sized) at(ses=( $\$min$   $\$minusd$   $\$mean$   $\$plusd$   $\$max$ )) atmeans
```

Conditional marginal effects

Number of obs = 7,185

Expression: Linear prediction, fixed portion, predict()

dy/dx wrt: 0.sized 1.sized 2.sized

```
1._at: ses = -3.758
      0.sized = .2499652 (mean)
      1.sized = .2455115 (mean)
      2.sized = .2509395 (mean)
      3.sized = .2535839 (mean)
2._at: ses = -.7792118
      0.sized = .2499652 (mean)
      1.sized = .2455115 (mean)
      2.sized = .2509395 (mean)
      3.sized = .2535839 (mean)
3._at: ses = .0001434
      0.sized = .2499652 (mean)
      1.sized = .2455115 (mean)
      2.sized = .2509395 (mean)
      3.sized = .2535839 (mean)
4._at: ses = .7794985
      0.sized = .2499652 (mean)
      1.sized = .2455115 (mean)
      2.sized = .2509395 (mean)
      3.sized = .2535839 (mean)
5._at: ses = 2.692
      0.sized = .2499652 (mean)
      1.sized = .2455115 (mean)
      2.sized = .2509395 (mean)
      3.sized = .2535839 (mean)
```

| Delta-method

	dy/dx	std. err.	z	P> z	[95% conf. interval]	

0.sized						
_at						
1	3.249233	1.296846	2.51	0.012	.7074619	5.791005
2	.8328605	.5674189	1.47	0.142	-.2792601	1.944981
3	.2006529	.5149248	0.39	0.697	-.8085812	1.209887
4	-.4315548	.5764797	-0.75	0.454	-1.561434	.6983247
5	-1.982963	1.011277	-1.96	0.050	-3.96503	-.0008955

1.sized						
_at						
1	4.625166	1.311975	3.53	0.000	2.053742	7.196589
2	1.962812	.5725967	3.43	0.001	.8405435	3.085081
3	1.266248	.5096593	2.48	0.013	.2673338	2.265161
4	.5696828	.5608856	1.02	0.310	-.5296327	1.668998
5	-1.139655	.9850071	-1.16	0.247	-3.070233	.7909237

2.sized						
_at						
1	3.000998	1.308259	2.29	0.022	.4368563	5.565139
2	1.528898	.5730777	2.67	0.008	.405686	2.652109
3	1.143745	.5141005	2.22	0.026	.1361262	2.151363
4	.7585918	.5689256	1.33	0.182	-.3564819	1.873665
5	-.1865555	.996368	-0.19	0.851	-2.139401	1.76629

3.sized	(base outcome)					

Note: dy/dx for factor levels is the discrete change from the base level.

Compared to largest schools, the smallest schools have higher math achievement scores for students with minimum SES (3.2 units higher); there are no significant differences between the largest and smallest schools at mean SES, maximum SES, or at mean+1SD and mean-1SD. The second and third category of school size are different from the largest schools in terms of math achievement for students whose SES is at mean or below (including mean itself, mean-1SD, and minimum SES). For minimum SES, the gap between the largest schools and second smallest is 4.6; the gap between the largest schools and second largest is 3 units. At mean SES, those gaps are 1.27 and 1.14, respectively. If we wanted to see those comparisons to another SIZE category, we could reestimate the model and the margins:

```
. mixed mathach c.ses##ib2.sized || id:ses, cov(unstr)
```

```
Performing EM optimization ...
```

```
Performing gradient-based optimization:
```

```
Iteration 0: log likelihood = -23309.269
```

```
Iteration 1: log likelihood = -23308.899
```

```
Iteration 2: log likelihood = -23308.899
```

```
Computing standard errors ...
```

```
Mixed-effects ML regression
Group variable: id
```

```
Number of obs      =      7,185
Number of groups   =       160
Obs per group:
    min =          14
    avg =          44.9
    max =           67
```

```
Log likelihood = -23308.899
```

```
Wald chi2(7)      =      457.54
Prob > chi2       =      0.0000
```

mathach	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
ses	2.417146	.2354425	10.27	0.000	1.955687	2.878605
sized						
0	-.9430464	.5394171	-1.75	0.080	-2.000284	.1141917
1	.1225602	.534395	0.23	0.819	-.9248346	1.169955
3	-1.143816	.514101	-2.22	0.026	-2.151435	-.1361962
sized#c.ses						
0	-.3169989	.3335864	-0.95	0.342	-.9708163	.3368184
1	-.3995763	.3322965	-1.20	0.229	-1.050865	.2517128
3	.4941943	.3188418	1.55	0.121	-.1307242	1.119113
_cons	13.19763	.38087	34.65	0.000	12.45114	13.94412

Random-effects parameters	Estimate	Std. err.	[95% conf. interval]	
id: Unstructured				
var(ses)	.3036591	.2180771	.0743162	1.240763
var(_cons)	4.534295	.635098	3.445767	5.966693
cov(ses,_cons)	-.0604166	.2803227	-.6098389	.4890057
var(Residual)	36.80666	.6285147	35.59518	38.05937

LR test vs. linear model: chi2(3) = 428.23 Prob > chi2 = 0.0000

Note: LR test is conservative and provided only for reference.

. margins, dydx(sized) at(ses=(\$min \$minusd \$mean \$plusd \$max)) atmeans

Conditional marginal effects Number of obs = 7,185

Expression: Linear prediction, fixed portion, predict()

dy/dx wrt: 0.sized 1.sized 3.sized

```

1._at: ses      =   -3.758
      0.sized =   .2499652 (mean)
      1.sized =   .2455115 (mean)
      2.sized =   .2509395 (mean)
      3.sized =   .2535839 (mean)
2._at: ses      =  -0.7792118
      0.sized =   .2499652 (mean)
      1.sized =   .2455115 (mean)
      2.sized =   .2509395 (mean)
      3.sized =   .2535839 (mean)
3._at: ses      =   .0001434
      0.sized =   .2499652 (mean)
      1.sized =   .2455115 (mean)
      2.sized =   .2509395 (mean)
      3.sized =   .2535839 (mean)
4._at: ses      =   .7794985
      0.sized =   .2499652 (mean)
      1.sized =   .2455115 (mean)
      2.sized =   .2509395 (mean)
      3.sized =   .2535839 (mean)
5._at: ses      =    2.692
      0.sized =   .2499652 (mean)
      1.sized =   .2455115 (mean)
      2.sized =   .2509395 (mean)
      3.sized =   .2535839 (mean)

```

		Delta-method		z	P> z	[95% conf. interval]	
		dy/dx	std. err.				

0.sized							
	_at						
	1	.2482356	1.37469	0.18	0.857	-2.446108	2.942579
	2	-.6960371	.6034782	-1.15	0.249	-1.878833	.4867584
	3	-.9430918	.5394161	-1.75	0.080	-2.000328	.1141444
	4	-1.190147	.5940835	-2.00	0.045	-2.354529	-.0257643
	5	-1.796407	1.038212	-1.73	0.084	-3.831266	.2384509

1.sized							
	_at						
	1	1.624168	1.388971	1.17	0.242	-1.098166	4.346502
	2	.4339148	.6083492	0.71	0.476	-.7584276	1.626257
	3	.1225029	.5343919	0.23	0.819	-.924886	1.169892
	4	-.188909	.5789637	-0.33	0.744	-1.323657	.945839
	5	-.9530993	1.012641	-0.94	0.347	-2.937839	1.03164

2.sized		(base outcome)					

3.sized							
	_at						
	1	-3.000998	1.308259	-2.29	0.022	-5.565139	-.4368563
	2	-1.528898	.5730777	-2.67	0.008	-2.652109	-.405686
	3	-1.143745	.5141005	-2.22	0.026	-2.151363	-.1361262
	4	-.7585918	.5689256	-1.33	0.182	-1.873665	.3564819
	5	.1865555	.996368	0.19	0.851	-1.76629	2.139401

Note: dy/dx for factor levels is the discrete change from the base level.

Example 4: Two continuous variables

Here, we will examine an interaction of SES and SIZE as a continuous variable rather than a set of dummies. We may want to use mean-centered SIZE variable here (we created it above).

```
. mixed mathach c.ses##c.sizem || id: ses, cov(unstr)
```

```
Mixed-effects ML regression          Number of obs    =      7,185
Group variable: id                   Number of groups =      160
                                      Obs per group:
                                      min =          14
                                      avg =          44.9
                                      max =           67
                                      Wald chi2(3)      =      438.94
                                      Prob > chi2       =      0.0000
Log likelihood = -23314.32
```

mathach	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
ses	2.392577	.1153567	20.74	0.000	2.166482	2.618672
sizem	-.000266	.0003041	-0.87	0.382	-.0008621	.0003301
c.ses#c.sizem	.0004921	.0001853	2.65	0.008	.0001288	.0008554
_cons	12.67515	.1890736	67.04	0.000	12.30458	13.04573

Random-effects parameters	Estimate	Std. err.	[95% conf. interval]	

```

id: Unstructured |
      var(ses) | .3168798 .221958 .0802919 1.250597
      var(_cons) | 4.788075 .6640414 3.64847 6.283637
      cov(ses,_cons) | -.1128863 .2887151 -.6787576 .4529849
-----+-----
      var(Residual) | 36.82078 .628957 35.60845 38.07438
-----+-----
LR test vs. linear model: chi2(3) = 461.36 Prob > chi2 = 0.0000

```

Note: LR test is conservative and provided only for reference.

SES as the focal variable, SIZE as the moderator:

The main effect of SES shows that at average school size, SES has a positive effect on math achievement – in average size schools, one unit increase in SES translates into 2.4 units increase in math achievement. That effect gets more pronounced in schools that are larger in size. Again, the unit for school size here is one student so to understand how much that effect is moderated, it is better to look at a standard deviation of size:

```

. sum sizem

```

Variable	Obs	Mean	Std. Dev.	Min	Max
sizem	7,185	-40.96321	604.1725	-997.825	1615.175

```

. global sizesd=r(sd)
. qui mixed mathach c.ses##c.sizem || id: ses, cov(unstr)
. mat list e(b)
e(b) [1,8]
      mathach:      mathach:      mathach:      mathach:      lns1_1_1:
              ses              sizem              c.sizem              _cons              _cons
y1      2.3925766      -.00026597      .00049209      12.675153      -.57461631

      lns1_1_2:      atr1_1_1_2:      lnsig_e:
              _cons              _cons              _cons
y1      .78306421      -.0919039      1.8030312

. di e(b) [1,1]+e(b) [1,3]*$sizesd
2.6898841

```

So for a school that's one SD above the average size, one unit increase in SES translates into 2.7 units increase in math achievement. And for a school that's one SD below the average size, one unit increase in SES translates into only 2.1 units increase in math achievement, as per this calculation:

```

. di e(b) [1,1]-e(b) [1,3]*$sizesd
2.0952691

```

Since SES is continuous, we may want to show these differences graphically. For predicted values and graphs, it is better to use uncentered size, even though we would present the coefficients from the centered model.

```

. mixed mathach c.ses##c.size || id: ses, cov(unstr)

```

```

Mixed-effects ML regression
Group variable: id
Number of obs = 7,185
Number of groups = 160
Obs per group:
    min = 14
    avg = 44.9
    max = 67
Wald chi2(3) = 438.94
Prob > chi2 = 0.0000
Log likelihood = -23314.32

```

	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
ses	1.852348	.2332827	7.94	0.000	1.395122	2.309573
size	-.000266	.0003041	-0.87	0.382	-.0008621	.0003301
c.ses#c.size	.0004921	.0001853	2.65	0.008	.0001288	.0008554
_cons	12.96714	.3817026	33.97	0.000	12.21902	13.71526

Random-effects parameters	Estimate	Std. err.	[95% conf. interval]	
id: Unstructured				
var(ses)	.3168798	.221958	.0802919	1.250597
var(_cons)	4.788075	.6640414	3.64847	6.283637
cov(ses,_cons)	-.1128863	.2887151	-.6787576	.4529849
var(Residual)	36.82078	.628957	35.60845	38.07438

```

LR test vs. linear model: chi2(3) = 461.36
Prob > chi2 = 0.0000

```

Note: LR test is conservative and provided only for reference.

```
. sum size if tagged==1
```

Variable	Obs	Mean	Std. dev.	Min	Max
size	160	1097.825	629.5064	100	2713

```
. global sizemin=r(min)
```

```
. global sizemax=r(max)
```

```
. global sizemean=r(mean)
```

```
. global sizeplussd=r(mean)+r(sd)
```

```
. global sizeminussd=r(mean)-r(sd)
```

```
. margins, dydx(ses) at(size=($sizemin $sizeminussd $sizemean $sizeplussd $sizemax))
atmeans
```

```
Conditional marginal effects
```

```
Number of obs = 7,185
```

```
Expression: Linear prediction, fixed portion, predict()
```

```
dy/dx wrt: ses
```

```
1._at: ses = .0001434 (mean)
      size = 100
```

```
2._at: ses = .0001434 (mean)
      size = 468.3186
```

```
3._at: ses = .0001434 (mean)
      size = 1097.825
```

```
4._at: ses = .0001434 (mean)
```



```

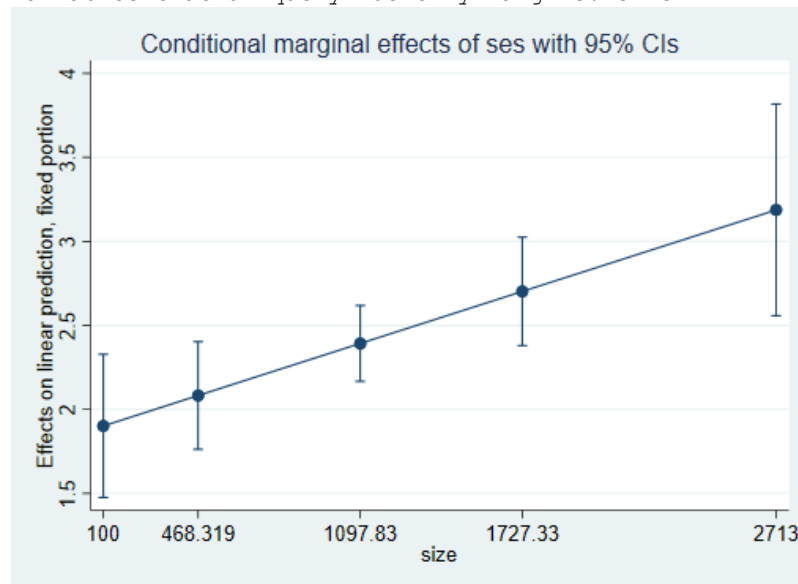
size = 1727.331
5._at: ses = .0001434 (mean)
size = 2713

```

		Delta-method		z	P> z	[95% conf. interval]	
		dy/dx	std. err.				
ses	_at						
	1	1.901557	.2173656	8.75	0.000	1.475528	2.327585
	2	2.082803	.1635671	12.73	0.000	1.762217	2.403388
	3	2.392577	.1153567	20.74	0.000	2.166482	2.618672
	4	2.702351	.1645847	16.42	0.000	2.379771	3.024931
	5	3.187389	.3214957	9.91	0.000	2.557269	3.817509

```
. marginsplot
```

```
Variables that uniquely identify margins: size
```



This shows the entire range of effect sizes for SES depending on size. We can see those as slopes if we calculate predicted values and plot them, but for that, let's also save the minimum, maximum etc. for SES:

```
. sum ses
```

Variable	Obs	Mean	Std. dev.	Min	Max
ses	7,185	.0001434	.7793552	-3.758	2.692

```
. global sesmin=r(min)
```

```
. global sesmax=r(max)
```

```
. global sesmean=r(mean)
```

```
. global sesplussd=r(mean)+r(sd)
```

```
. global sesminusd=r(mean)-r(sd)
```

```
. margins, at(size=($sizemin $sizeminusd $sizemean $sizeplussd $sizemax) ses=($sesmin $sesminusd $sesmean $sesplussd $sesmax)) atmeans
```

Adjusted predictions

Number of obs = 7,185

Expression: Linear prediction, fixed portion, predict()

```
1._at: ses = -3.758
      size = 100
2._at: ses = -3.758
      size = 468.3186
3._at: ses = -3.758
      size = 1097.825
4._at: ses = -3.758
      size = 1727.331
5._at: ses = -3.758
      size = 2713
6._at: ses = -.7792118
      size = 100
7._at: ses = -.7792118
      size = 468.3186
8._at: ses = -.7792118
      size = 1097.825
9._at: ses = -.7792118
      size = 1727.331
10._at: ses = -.7792118
      size = 2713
11._at: ses = .0001434
      size = 100
12._at: ses = .0001434
      size = 468.3186
13._at: ses = .0001434
      size = 1097.825
14._at: ses = .0001434
      size = 1727.331
15._at: ses = .0001434
      size = 2713
16._at: ses = .7794985
      size = 100
17._at: ses = .7794985
      size = 468.3186
18._at: ses = .7794985
      size = 1097.825
19._at: ses = .7794985
      size = 1727.331
20._at: ses = .7794985
      size = 2713
21._at: ses = 2.692
      size = 100
22._at: ses = 2.692
      size = 468.3186
23._at: ses = 2.692
      size = 1097.825
24._at: ses = 2.692
      size = 1727.331
25._at: ses = 2.692
      size = 2713
```

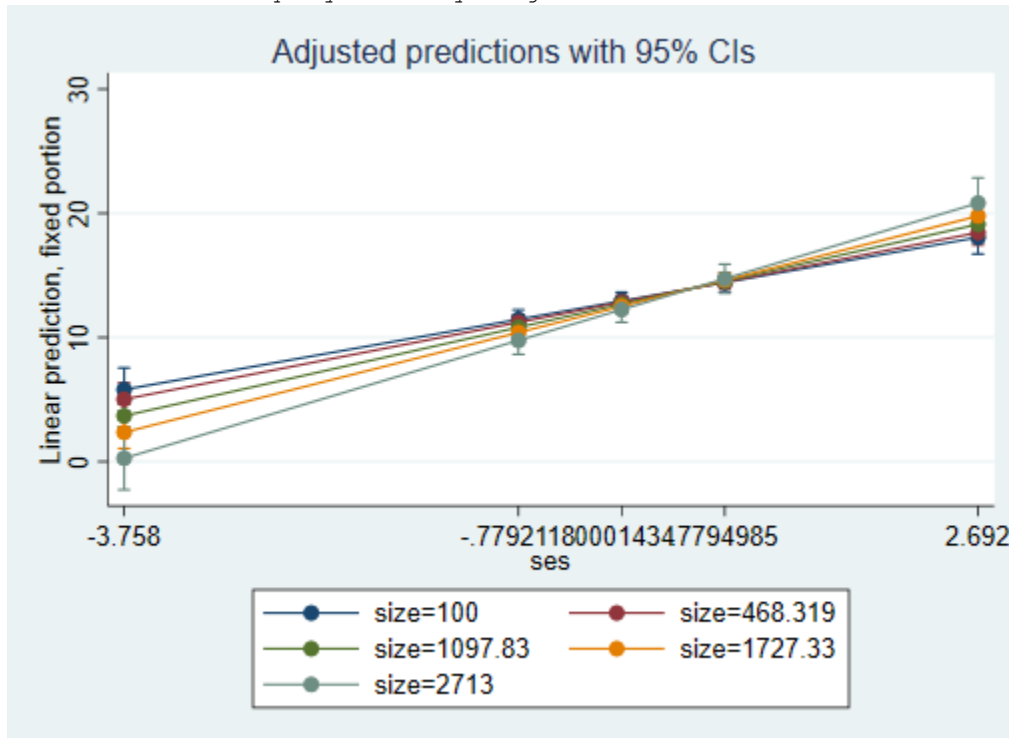
	Margin	Delta-method std. err.	z	P> z	[95% conf. interval]	

_at						
1	5.794493	.8937958	6.48	0.000	4.042685	7.5463

2		5.01541	.676854	7.41	0.000	3.6888	6.342019
3		3.683851	.4786823	7.70	0.000	2.745651	4.622051
4		2.352292	.6678443	3.52	0.000	1.043341	3.661243
5		.2673639	1.296626	0.21	0.837	-2.273976	2.808704
6		11.45883	.395233	28.99	0.000	10.68418	12.23347
7		11.21964	.2993116	37.48	0.000	10.633	11.80628
8		10.81083	.2120376	50.99	0.000	10.39524	11.22642
9		10.40202	.2961318	35.13	0.000	9.821614	10.98243
10		9.761919	.5745359	16.99	0.000	8.63585	10.88799
11		12.94081	.3556007	36.39	0.000	12.24385	13.63778
12		12.84288	.2674408	48.02	0.000	12.31871	13.36705
13		12.6755	.1890731	67.04	0.000	12.30492	13.04607
14		12.50811	.2707095	46.20	0.000	11.97753	13.03869
15		12.24603	.5285042	23.17	0.000	11.21018	13.28188
16		14.4228	.3925439	36.74	0.000	13.65343	15.19217
17		14.46612	.2931934	49.34	0.000	13.89147	15.04077
18		14.54016	.2066463	70.36	0.000	14.13514	14.94518
19		14.6142	.3029515	48.24	0.000	14.02043	15.20798
20		14.73013	.5950645	24.75	0.000	13.56383	15.89644
21		18.05953	.6820231	26.48	0.000	16.72279	19.39627
22		18.44949	.5090456	36.24	0.000	17.45177	19.4472
23		19.11597	.3581556	53.37	0.000	18.414	19.81794
24		19.78245	.525949	37.61	0.000	18.75161	20.81329
25		20.82602	1.03428	20.14	0.000	18.79887	22.85317

```
-----
. marginsplot, x(ses)
```

Variables that uniquely identify margins: size ses



This is a bit too busy, let's simplify to look at two extremes, minimum size and maximum size:

```
. margins, at(size=($sizemin $sizemax) ses=($sesmin $sesmean $sesmax)) atmeans
Adjusted predictions          Number of obs      =          7,185
```

```

Expression : Linear prediction, fixed portion, predict()
1._at     : ses           =    -3.758
           : size         =     100

2._at     : ses           =    -3.758
           : size         =    2713

3._at     : ses           =     .0001434
           : size         =     100

4._at     : ses           =     .0001434
           : size         =    2713

5._at     : ses           =     2.692
           : size         =     100

6._at     : ses           =     2.692
           : size         =    2713

```

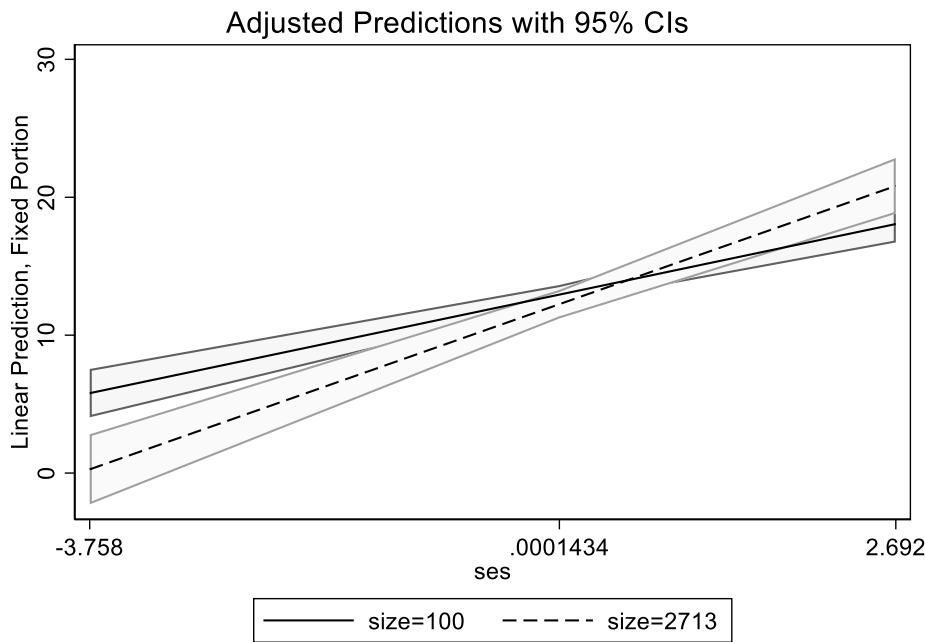
	Margin	Delta-method Std. Err.	z	P> z	[95% Conf. Interval]	
_at						
1	5.794493	.8937958	6.48	0.000	4.042685	7.5463
2	.2673639	1.296626	0.21	0.837	-2.273976	2.808704
3	12.94081	.3556007	36.39	0.000	12.24385	13.63778
4	12.24603	.5285042	23.17	0.000	11.21018	13.28188
5	18.05953	.6820231	26.48	0.000	16.72279	19.39627
6	20.82602	1.03428	20.14	0.000	18.79887	22.85317

```

. marginsplot, x(ses) plotop(msymbol(i)) recastci(rarea) ciopts(fintensity(5))
scheme(slmono)

```

Variables that uniquely identify margins: size ses



SIZE as the focal variable, SES as the moderator:

For a student with mean SES, when school size increases by 1 student, math achievement does not change – the main effect of school size is $-.000266$ but it's not statistically significant. Let's use margins to see what that effect of size looks like at other values of SES:

```
. margins, dydx(size) at(ses=($sesmin $sesminussd $sesmean $sesplussd $sesmax))
atmeans
```

Conditional marginal effects

Number of obs = 7,185

Expression: Linear prediction, fixed portion, predict()

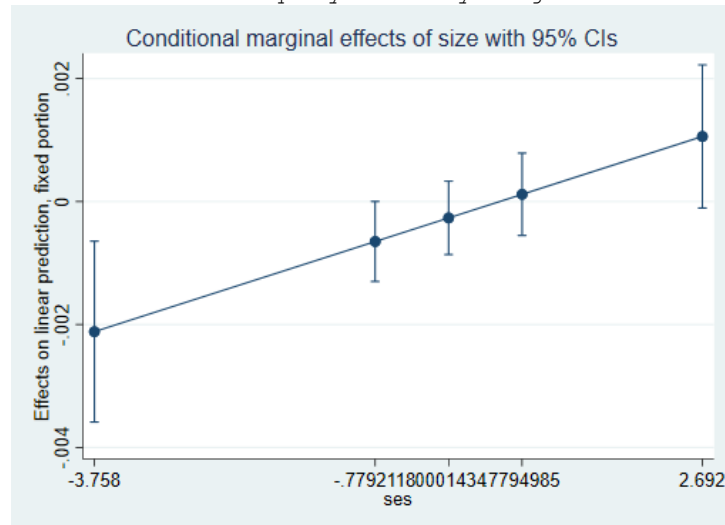
```
dy/dx wrt: size
1._at: ses = -3.758
      size = 1056.862 (mean)
2._at: ses = -.7792118
      size = 1056.862 (mean)
3._at: ses = .0001434
      size = 1056.862 (mean)
4._at: ses = .7794985
      size = 1056.862 (mean)
5._at: ses = 2.692
      size = 1056.862 (mean)
```

		Delta-method				
		dy/dx	std. err.	z	P> z	[95% conf. interval]
size						
	_at					
	1	-.0021152	.0007501	-2.82	0.005	-.0035853 - .0006452
	2	-.0006494	.000332	-1.96	0.050	-.0013001 1.31e-06
	3	-.0002659	.0003041	-0.87	0.382	-.000862 .0003302
	4	.0001176	.0003413	0.34	0.730	-.0005514 .0007866
	5	.0010587	.0005935	1.78	0.074	-.0001046 .002222

Here we can see that the effect of school size is only significant for those with SES lower than 1 SD below the mean (the p value is exactly .05 for 1 SD below the mean, so we assume it will be significant anywhere below that). Let's see that difference in effect sizes graphically:

```
. marginsplot
```

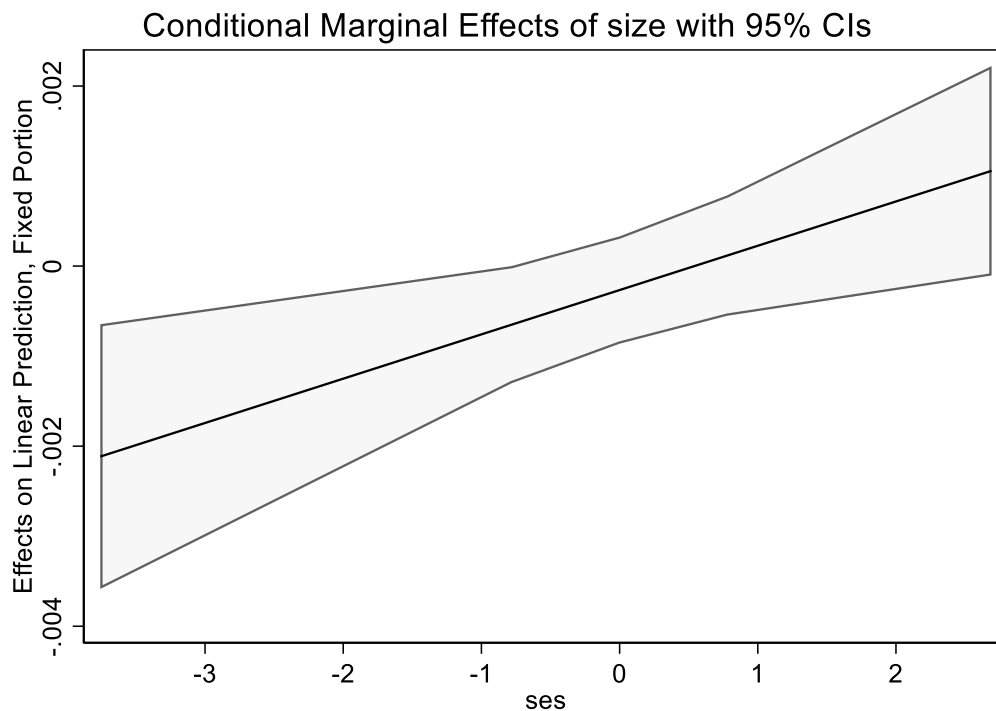
Variables that uniquely identify margins: ses



We can see which effects are statistically significant based on whether 0 is within the confidence interval or not. That might be better visible with a confidence band:

```
. marginsplot, x(ses) plotop(msymbol(i)) recastci(rarea) ciopts(fintensity(5))
> scheme(slmono) xlabel(-3 -2 -1 0 1 2)
```

Variables that uniquely identify margins: ses



Finally, let's see those effects of size as actual slopes:

```
. margins, at(size=($sizemin $sizeminusd $sizemean $sizeplusd $sizemax) ses=($sesmin $sesminusd $sesmean $sesplusd $sesmax)) atmeans
```

Adjusted predictions

Number of obs = 7,185

Expression: Linear prediction, fixed portion, predict()

- 1._at: ses = -3.758
size = 100
- 2._at: ses = -3.758
size = 468.3186
- 3._at: ses = -3.758
size = 1097.825
- 4._at: ses = -3.758
size = 1727.331
- 5._at: ses = -3.758
size = 2713
- 6._at: ses = -.7792118
size = 100
- 7._at: ses = -.7792118
size = 468.3186
- 8._at: ses = -.7792118
size = 1097.825
- 9._at: ses = -.7792118

```

size = 1727.331
10._at: ses = -.7792118
size = 2713
11._at: ses = .0001434
size = 100
12._at: ses = .0001434
size = 468.3186
13._at: ses = .0001434
size = 1097.825
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15._at: ses = .0001434
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16._at: ses = .7794985
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17._at: ses = .7794985
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18._at: ses = .7794985
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20._at: ses = .7794985
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21._at: ses = 2.692
size = 100
22._at: ses = 2.692
size = 468.3186
23._at: ses = 2.692
size = 1097.825
24._at: ses = 2.692
size = 1727.331
25._at: ses = 2.692
size = 2713

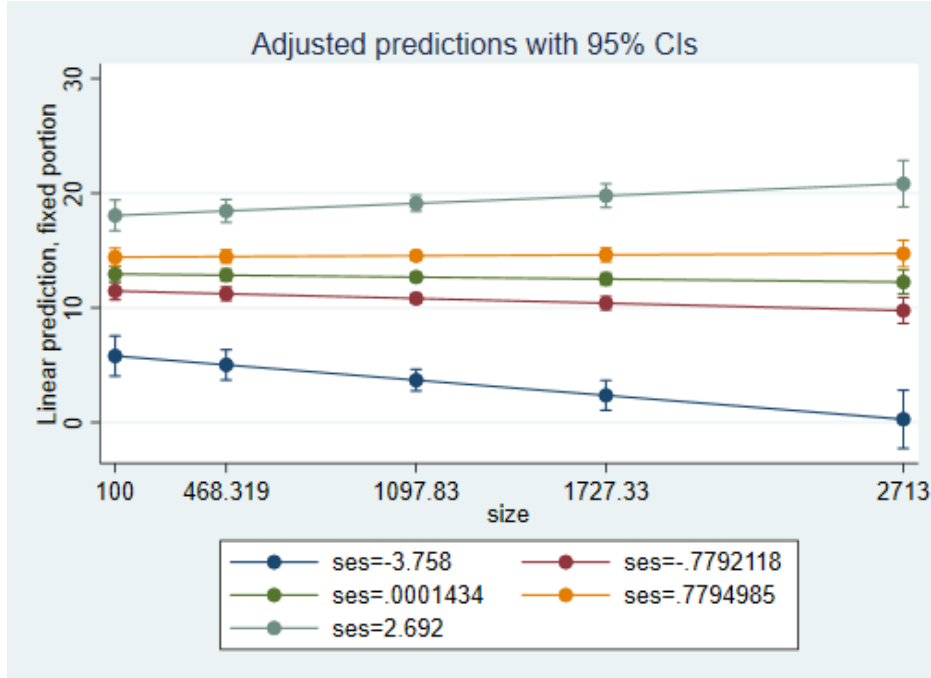
```

	Margin	Delta-method std. err.	z	P> z	[95% conf. interval]	
_at						
1	5.794493	.8937958	6.48	0.000	4.042685	7.5463
2	5.01541	.676854	7.41	0.000	3.6888	6.342019
3	3.683851	.4786823	7.70	0.000	2.745651	4.622051
4	2.352292	.6678443	3.52	0.000	1.043341	3.661243
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9	10.40202	.2961318	35.13	0.000	9.821614	10.98243
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11	12.94081	.3556007	36.39	0.000	12.24385	13.63778
12	12.84288	.2674408	48.02	0.000	12.31871	13.36705
13	12.6755	.1890731	67.04	0.000	12.30492	13.04607
14	12.50811	.2707095	46.20	0.000	11.97753	13.03869
15	12.24603	.5285042	23.17	0.000	11.21018	13.28188
16	14.4228	.3925439	36.74	0.000	13.65343	15.19217
17	14.46612	.2931934	49.34	0.000	13.89147	15.04077
18	14.54016	.2066463	70.36	0.000	14.13514	14.94518
19	14.6142	.3029515	48.24	0.000	14.02043	15.20798
20	14.73013	.5950645	24.75	0.000	13.56383	15.89644
21	18.05953	.6820231	26.48	0.000	16.72279	19.39627
22	18.44949	.5090456	36.24	0.000	17.45177	19.4472
23	19.11597	.3581556	53.37	0.000	18.414	19.81794
24	19.78245	.525949	37.61	0.000	18.75161	20.81329

25 | 20.82602 1.03428 20.14 0.000 18.79887 22.85317

```
. marginsplot, x(size)
```

Variables that uniquely identify margins: size ses



```
. marginsplot, x(size) plotop(msymbol(i)) recastci(rarea) ciopts(fintensity(5))  
scheme(slmono)
```

Variables that uniquely identify margins: size ses

